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Re-examining sustained attention deficits as vulnerability indicators for schizophrenia: Stability in the long term course

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Abstract

This study examined the longitudinal patterns in the sustained attention deficits detected by the Continuous Performance Test (CPT) and the factors influencing such changes in consecutively admitted schizophrenia patients (n = 224) followed up for 4–7 year. Exploratory growth mixture modeling analyses of subjects' CPT performances over successive follow-ups revealed that three major (accounting for 92.8%) plus one minor subgroups could be delineated. Subgrouping was then performed on a subsample of 104 subjects who had at least 3 times of CPT data. Based on subjects' adjusted z score of the test sensitivity index d' derived from comparing with a community sample, patients were divided into three subgroups: no impairment (≥ -1), moderate impairment (-2.5 to -1), and severe impairment (<-2.5). The trajectory taken by individual patient was analyzed according to the initial subgroup status and subsequent changes, controlling for relevant basic and clinical characteristics. Both growth mixture modeling and subgroup status analyses found that around one third of those with severe impairment at baseline showed persistent severe impairment. Those with no impairment were stable and exhibited least tendency for further performance deterioration. Those with moderate impairment tended to fluctuate markedly, mainly towards the better rather than the worse. Previous subgrouping status and concurrent task-taking strategy predicted the performance subgroup status at follow-ups, while clinical symptoms and disease course factors did not. We concluded that there is substantial heterogeneity in schizophrenia patients' long term pattern in sustained attention deficits and those with severe impairment might represent a subgroup with stable vulnerability to schizophrenia.

Keywords: Schizophrenia; Continuous performance test; Sustained attention; Longitudinal study; Growth mixture modeling; Vulnerability indicator

1. Introduction

Sustained attention deficits measured by the Continuous Performance Test (CPT) (Rosvold et al., 1956) have been considered as potential vulnerability indicators for schizophrenia (Chen and Faraone, 2000; Cornblatt and Keilp, 1994; Keri and Janka, 2004). They were demonstrated in schizophrenic patients across illness stages (Asarnow and

MacCrimmon, 1981; Nuechterlein and Dawson, 1984) and remained relatively unchanged despite fluctuations in clinical symptoms, in contrast to those manifested by affective disorders (Liu et al., 2002; Nuechterlein et al., 1994). In addition, the performance deficits were not amendable to treatment with neuroleptics when more difficult CPT versions were used (Finkelstein et al., 1997; Hong et al., 2002; Liu et al., 2000). People at high risk of schizophrenia, such as relatives of schizophrenia patients, were found to have worse CPT scores than normal controls (Egan et al., 2000; Mirsky et al., 1995; Saoud et al., 2000) and the CPT deficits were associated with both clinical and psychometric schizotypes (Chen et al., 1997; Laurent et al.,

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2000; Roitman et al., 1997). The differential longitudinal changes in CPT performance thus might be indicative of the nature of the underlying vulnerability. For those schizophrenia patients in whom the genetic vulnerability plays a major role, their sustained attention impairments will be more severe and stable in the long-term course; otherwise, the severity of the impairments will be determined more by state-dependent factors, such as clinical states and treatment-related side effects, hence tend to fluctuate over the course (Chen and Faraone, 2000; Nuechterlein et al., 1994).

Despite such important theoretical implication, the longitudinal stability of CPT deficits in schizophrenic patients over the long-term course was seldom examined. Past follow-up studies were limited mainly to weeks or months after the immediate postpsychotic phase, rendering it premature to extrapolate the results to longer period (Epstein et al., 1996; Finkelstein et al., 1997; Hong et al., 2002; Liu et al., 2000; Nuechterlein et al., 1992). Indeed, a few studies with follow-up durations beyond one year did report improvements in the CPT deficits, despite their limited sample size (Cosway et al., 2002; Olbrich et al., 2001). However, most of the previous studies did not control for possible confounding factors of CPT performances, such as demographic characteristics, clinical symptom, medication, and treatment-related side effects. In addition, past follow-up studies approached the stability issues solely by analyzing the changes in the group means across different time points (Cosway et al., 2002; Epstein et al., 1996; Finkelstein et al., 1997; Hong et al., 2002; Liu et al., 2000; Nuechterlein et al., 1992; Olbrich et al., 2001), which might have masked the individual differences in the trajectories.

In this study, two major analytic strategies were undertaken to examine the issue of longitudinal stability in the CPT performance of schizophrenia patients with 4–7 years of follow-up. Analysis 1 was to explore whether there were indeed subgroups within schizophrenia through data-driven growth mixture modeling, hence to provide information about the optimal number of the subgroups as well as the characteristics of each subgroup's trajectories. Analysis 2 was to testify the hypothesis that CPT performance subgroups categorized according to their baseline CPT performance had different longitudinal stability over the follow-up periods. The longitudinal stability was defined in terms of the tendencies of the subjects to remain in their original subgroup status and was further analyzed with control for potential confounders to disentangle the effects of state-dependent factors from those of the vulnerability indicators.

2. Methods

2.1. Subjects

The study subjects were schizophrenic patients participating in two serial prospective follow-up studies for studying longitudinal course of psychopathological and

neuropsychological manifestations of schizophrenia in Taiwan. The first study spanned from July 1993 to June 1998 and the second study was an extension of the first with the same study design from July 1998 to December 2001 (Chang et al., 2002; Hwu et al., 2002). Both were hitherto referred to as the Taiwan Psychopathology Study of Schizophrenia. During the enrolling period, one university and two university-affiliated hospitals recruited 234 consecutively admitted schizophrenia inpatients after written informed consent was obtained. The diagnosis was made according to the DSM-IV diagnostic criteria of schizophrenia tentatively at admission by senior attending psychiatrists. In addition, interview in person using the Chinese version of the Diagnostic Interview for Genetic Studies (DIGS) (Chen et al., 1998b; Nurnberger et al., 1994) by trained research assistants was used to collect all available information, such as care-giver report and previous medical records from other hospitals. The diagnoses were subsequently confirmed at discharge by independent assessments of three senior psychiatrists, based on catamnestic data and psychopathological manifestations during index admission. Among the 234 subjects, 6 (2.5%) had inconsistent diagnoses and final diagnoses were reached through consensus meeting. Exclusion criteria were a history of electroconvulsive therapy in previous 6 months, mental retardation, trauma-related loss in consciousness for more than 30 min, psychoactive substance abuse, or physical illness that might cast doubt on the diagnosis. The studies were approved by the Institutional Review Boards of the participating hospitals.

Although the main interest of this study was the changes within individual schizophrenic patients, we were still keen to know the magnitude of impairments in these patients compared to normal subjects. Hence a representative sample of community subjects was included as the norm. The normal comparison subjects were systematically sampled from the 1993 and 1994 voter lists of Chinshan Township, north of Taipei (Chen et al., 1998a). A total of 345 subjects, about 65% of selected local residents, were successfully tested. Subjects who had a diagnosis of psychosis according to the outpatient records at Chinshan Health Station or stroke were excluded from the normative data.

2.2. Measures

2.2.1. Clinical assessments and psychopathology measures

Baseline information regarding age at onset, duration of illness, previous medication and hospitalization histories were collected systemically. The Chinese version of the Positive and Negative Syndrome Schedules (PANSS) (Kay et al., 1987), which was found to have sufficient inter-rater reliability (Cheng et al., 1996) was used by trained senior research psychiatrists to assess clinical psychopathology with the following time frame: admission, discharge, 3 months, 6 months, 12 months, and then on a yearly basis.

Our previous study on the symptom dimensions of the PANSS revealed a five-factor structure (Hwu et al., 2002) and the compositions of these factors were largely in accordance with those from other studies (Mass et al., 2000). As the symptom dimensions were more closely related to neurocognitive performances than the original PANSS sub-scales (Liu et al., 1997), we averaged the summed-up individual PANSS item ratings within the same factor to yield five mean factor scores, i.e., Negative, Disorganization, Positive, Depression, and Excitement. Subsequent analyses on clinical psychopathology were all done on these mean factor scores.

2.2.2. CPT

The subjects were assessed with the CPT, version 2.20, operated on a Sunrise Systems machine (Pembroke, Mass.). The procedure and its test-retest reliability have been described in detail elsewhere (Chen et al., 1998a). Briefly, numbers from 0 to 9 were randomly presented for 50 msec each, at a rate of one per second for two sessions (undegraded and 25% degraded). A total of 331 trials, 34 (10%) of which were target stimuli (the number 9 preceded by the number 1), were presented over 5 min for each session. During the second session (25% degraded session), a pattern of snow was used to toggle background and foreground so that the image was visually distorted. Two signal-detection indices of CPT performance—sensitivity (d')and response criterion (natural logarithm of $\beta \lceil \ln \beta \rceil$ —were derived from the hit rate (probability of response to target trials) and false-alarm rate (probability of response to nontarget trials) (Davies and Parasuraman, 1982; Swets, 1973). The sensitivity index d' measures an individual's ability to discriminate target stimuli from nontarget stimuli, while the response criterion index $\ln \beta$ measures the amount of perceptual evidence that the person requires to identify a stimulus as a target, hence an indicator of patient's task-taking strategy. Variables extraneous to the intended comparison of sensitivity, such as differences in level of motivation or cooperativeness, would yield differences in $\ln \beta$ rather than in d'.

The baseline CPT assessments were administered after patients' clinical conditions stabilized enough. Of the initial 234 patients enrolled in the Taiwan Psychopathology Study of Schizophrenia, 187 completed the baseline CPT assessment. About 4 years later, CPT assessments were repeated on a yearly basis for three times. Overall, each subject was scheduled to have 4 times CPT assessments over the entire study period.

2.3. Data analysis

2.3.1. Analysis 1: Growth mixture modeling

To explore whether subgroups with distinct trajectories of changes in CPT performances could be delineated within current sample, growth mixture modeling was used to analyze data from those patients who had at least one CPT assessment during the entire follow-up period (224 out of

the 234 patients enrolled) using MPlus software (Muthen and Muthen, 2005). Subjects not included did not differ in basic demographic characteristics from those included for this analysis. The number of subgroups was not set a priori. Rather, models with successive numbers of subgroups were compared against each other, starting from a model with 2 subgroups. The model with most significant difference with its previous model of one subgroup less was selected, based on the likelihood ratio test with significance level set at p < 0.05. In addition, on the basis of Bayesian Information Criteria (BIC), the model with the lowest BIC was preferred (Muthen et al., 2002).

2.3.2. Analysis 2: Subgroup status change

We then subgrouped patients according to their baseline CPT performance to examine whether different subgroups had different longitudinal stability over the follow-up periods. Of the entire 234 patients, 104 patients completed more than 3 times CPT assessments, 80 completed one or two CPT assessments and 40 patients did not complete any CPT test. Only subjects with CPT performance data for at least three time points were included for this analysis. For those with 4 times of CPT assessments (n = 58), in addition to the baseline and the last assessments, one additional CPT assessment was chosen between the remaining two to make the CPT test intervals more evenly distributed. The mean interval between the baseline assessment and the first follow-up was 1213.9 (SD = 390.8) days, between the first and the second follow-up was 597.0 (SD = 182.1)days, and between the baseline and the second follow-up was 1807.1 (SD = 400.5) days.

To better understand their magnitude of deficits compared to normal comparisons, patients' CPT raw scores were standardized against the community norm with adjustment for demographic features to provide adjusted z scores (Chen et al., 1998a). Since age, sex, and educational level have been demonstrated to influence the CPT performance in the community sample, the predictive score of a subject was calculated by using the regression coefficients obtained from the regression of the scores on the three covariates among the 345 community subjects. The difference between the raw score and the predictive score was then standardized by the root mean error of the regression and was defined as the adjusted z score of the subject. The adjusted z scores thus provide an estimate of the magnitude of deviation from the community norm.

As our previous studies demonstrated that the degrees of CPT performance impairments in schizophrenic probands were related to the performances in unaffected relatives (Chen et al., 2004; Chen et al., 1998b), we chose the interface values of the three-component distribution of -1 and -2.5 (adjusted z score of d') from our previous admixture analysis on the CPT performances in both the non-psychotic relatives of schizophrenia patients and normal comparison subjects (Chen et al., unpublished data) as cut-off points. Intriguingly, these cut-off points were compatible with the results of growth mixture modeling.

Accordingly, patients with an adjusted z score of d' > -1 were categorized as the group with no impairment (NI), those with an adjusted z score < -2.5 as the group with severe impairment (SI), and those with an adjusted z score in between as the group with moderate impairment (MI).

In assessing the change in CPT performance grouping status across time, two measures of stability were calculated: prospective consistency and retrospective consistency. Prospective consistency equals to the proportion of individuals in a group at time 1 remaining in the same group at time 2. In contrast, retrospective consistency equals the proportion of individuals in a group at time 2 belonging to the same group at time 1.

The stability of the CPT-based subgrouping was further evaluated by exploring whether a later subgroup assignment could be predicted by a previous CPT performance. As the SI group is our major interest, we collapsed the NI and MI groups into a single group to contrast with the SI group and logistic regression analysis with the CPT group status (severe vs. non-severe impairment) at time 2 as dependent variable and previous group assignment at time 1 as the independent was conducted separately for the three different follow-up time frames. The extraneous factors that might have affected the CPT performances, including concurrent $\ln \beta$ at time 2, age at onset, duration of illness, interval between the two CPT assessments, and the 5 mean PANSS-CH symptom factor scores, were included as covariates. The Wald test was used to estimate the significance of the regression coefficients. The significance for all statistical test was set at a p value smaller than 0.05 without correction for multiple comparisons.

3. Results

The growth mixture modeling analysis of the 224 subjects' changes in undegraded CPT d' revealed that a 4-subgroup model best fitted the data (Fig. 1). The model with 4 subgroups had a better fit than the model with 3 subgroups (p = 0.037) for the likelihood ratio test; adjusted BICs = 1959.8 and 1975.1, respectively). The number of subjects (%) in each subgroup was 62 (27.6%), 17 (7.8%), 91 (40.7%), and 54 (24.0%) with the corresponding mean d's being 0.08, 3.27, 3.43, and 0.92, respectively, indicating that the 4 subgroups were distinguished by their trajectories of changes as well as their initial severity of impairments. Overall, the three main subgroups all showed a trend toward improvements in CPT d' during the successive follow-ups, whereas the smallest subgroup had a deteriorating course in CPT performance despite its initial good performance. Because the smallest subgroup comprised only 7.8% of total cases, the major trends for the trajectories of changes within the entire sample were better represented by the other 3 major subgroups. Similar results were obtained for the degraded CPT (data not shown).

The 104 subjects included for analysis 2 were of relatively young ages and medium disease chronicity with an even distribution in sex (Table 1). Although these 104

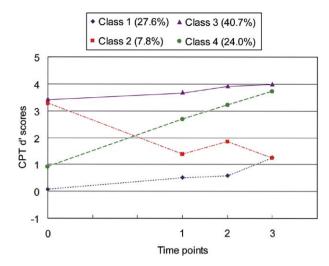


Fig. 1. Trajectories of the CPT performances of schizophrenia patients (n = 224). Time point 0 =baseline, 1 - 3 =first to third follow-up.

patients comprised only 44.4% out of the 234 patients enrolled, univariate analyses comparing them with those without CPT data (n=49) or those with less than three CPT assessments (n=80) did not reveal any significant differences in term of demographic characteristics, clinical course variables, and baseline symptom factor scores and CPT performances.

As the degraded CPT performance scores were highly correlated with the undegraded ones and the pattern of changes was similar to that of the undegraded ones too, only the results of the undegraded CPT were reported (Table 2). Patients' CPT performances as a group at the three assessments were all substantially impaired since their mean adjusted z scores of d' were significantly below those of the community sample. In contrast, the mean adjusted z scores of the $\ln \beta$ were not statistically deviant from the community norm. As a group, subjects' d' and $\ln \beta$ improved significantly across the follow-up period, no matter the raw scores or the adjusted z scores. The intraclass correlation coefficient for the three CPT assessments was 0.55 (F = 4.73 [df = 103, 206], p < 0.0001; 95.00% C.I.: 0.45-0.65), indicating that the CPT performances over the time points were moderately correlated. Antipsychotic dosage did not correlate with any CPT indices and patients receiving typical antipsychotics did not differ from those receiving atypical ones in their CPT performances.

At baseline, the CPT performance subgroups NI, MI and SI comprised 25%, 20%, and 55% of total subjects, respectively (Table 3). After 4–7 years, the corresponding fractions changed to 49%, 23%, and 28%, indicating an overall trend toward improvement in CPT performances. However, the consistency indices demonstrated the different trajectories of changes among the three groups. The NI subgroup showed least tendency to shift out, as revealed by their consistently high prospective consistencies across all the three time frames (58–90%), but substantial case shifting-in in the longer follow-ups, as indicated by the

Table 1
Demographic and clinical course characteristics, psychopathological manifestations, and CPT performance indices of patients recruited for the Taiwan Psychopathology Study of Schizophrenia, 1993–2001

Variable	Included in the subgrouping analyses ^a ($n = 104$)	Not included in the subgrouping analyses		
	N (%)	CPT test no. $\leq 2 (n = 80)$ N (%)	CPT not assessed $(n = 49)$ N (%)	
Male	53 (50.5)	41 (51.4)	23 (49.9)	
Being first admission	14 (13.3)	2 (4.1)	13 (16.5)	
Being on job in the previous one year	38 (36.1)	32 (40.0)	15 (30.6)	
	Mean (SD)	Mean (SD)	Mean (SD)	
Age	30.4 (7.0)	31.7 (8.2)	32.5 (7.4)	
Education (year)	11.2 (2.8)	10.9 (3.0)	10.7 (3.0)	
Age at onset	21.8 (6.2)	24.0 (7.1)	23.4 (6.1)	
Duration of illness (year)	8.6 (5.9)	7.6 (6.1)	9.1 (4.5)	
No. of previous admission	2.5 (2.1)	2.3 (2.1)	2.6 (2.5)	
Mean PANSS factor scores at baseline				
Negative	2.5 (0.8)	2.6 (0.9)	2.7 (0.9)	
Disorganization	1.9 (0.9)	2.2 (1.0)	2.0 (1.0)	
Positive	2.7 (1.0)	2.3 (1.1)	2.8 (1.1)	
Depression/anxiety	1.6 (0.8)	1.7 (0.8)	1.8 (0.7)	
Excitement	1.7 (1.0)	1.8 (1.3)	1.8 (1.0)	
Baseline undegraded CPT d'				
Raw score	2.1 (1.8)	1.8 (1.7)	_	
Adjusted z score	-3.2(2.5)	-3.6(2.3)	_	

All comparisons among the three groups were not statistically significant at the 0.05 level.

relatively low retrospective consistencies (39%). For the SI subgroup, the high retrospective consistencies across all three time frames (72–90%) indicated that few new cases moved into the SI subgroup. In contrast, the prospective consistencies of the SI subgroup decreased as the follow-up intervals lengthened. For the shorter interval from 1st follow-up to 2nd follow-up, 61% remained within the SI subgroup, while for the longest interval from baseline to 2nd follow-up, the percentage decreased remarkably to 39%. Thus, a substantial proportion of the initial SI subjects had improved performances at long-term follow-up. As to the MI subgroup, their prospective and retrospective consistencies across all time frames remained very low,

Table 2 Subjects' CPT performance indices and antipsychotic use at baseline and subsequent follow-ups (n=104)

	Baseline	1st follow-up	2nd follow-up
	Mean (SD)	Mean (SD)	Mean (SD)
Undegraded CPT indices			_
Raw score of d'^{a}	2.1 (1.8)	2.6 (1.7)	3.1 (1.6)
Raw score of $\ln \beta^{\rm b}$	1.4 (1.8)	1.8 (1.5)	2.1 (1.3)
Adjusted z score of d'^a	-3.2(2.5)	-2.4(2.4)	-1.7(2.1)
Adjusted z score of $\ln \beta^{b}$	-0.8(2.2)	-0.4(1.8)	0.0 (1.5)
Antipsychotic dosage ^c	902.3 (592.2)	826.6 (515.4)	815.6 (536.8)

^a p < 0.001, paired-t test, pairwise comparisons between successive assessments.

hence the most unstable among the three groups in terms of the case shifting tendency.

The determinants of later CPT performance subgroup assignment were explored by the logistic regression analyses across the three follow-up time frames (Table 4). Previous CPT grouping (by d' level) and concurrent CPT task-taking strategy ($\ln \beta$) were the major determinants for current CPT group assignment when factors extraneous to CPT were controlled for. Previous assignment in the SI subgroup as well as poorer concurrent $\ln \beta$ presented a higher risk to be classified as SI at follow-up (odds ratio ranged from 2.24 to 9.10 for previous CPT subgrouping and from 2.13 to 2.91 for $\ln \beta$, respectively). The diseaseassociated variables (duration of illness, age at onset, interval between follow-ups, and antipsychotic dosage) and clinical symptom dimensions (negative, disorganization, positive, depressive, and excitement factor) did not contribute significantly to later subgroup assignment. Adding other disease course characteristics, such as the number of rehospitalization and antipsychotic usage, did not alter the results.

4. Discussion

So far as we know, this is the longest follow-up study with multiple assessments of schizophrenia patients' sustained attention deficits, enabling us to address the issue of longitudinal stability beyond the immediate post psychotic phase. Overall, our results revealed a general tendency to improve in sustained attention during the entire follow-up period and about 1/3 of the schizophrenia

^a Patients receiving three CPT evaluations or more were included.

^b p < 0.05, paired-t test, pairwise comparisons between successive assessments, except comparison between 1st and 2nd follow-ups.

^c Chlorpromazine equivalents.

Table 3
Distribution of the undegraded CPT performance grouping during successive follow-ups

Initial grouping ^a	N (%)	Subsequent grouping			Consistency in grouping (%)	
		NI	MI	SI	Prospective	Retrospective
Baseline to 1st follow-up	104 (100)	38 (37)	23 (22)	43 (41)		
No impairment (NI)	26 (25)	15	7	4	57.6	39.5
Moderate impairment (MI)	21 (20)	10	3	8	14.3	13.0
Severe impairment (SI)	57 (55)	13	13	31	54.4	72.1
Baseline to 2nd follow-up	104 (100)	51 (49)	24 (23)	29 (28)		
No impairment	26 (25)	20	4	2	76.9	39.2
Moderate impairment	21 (20)	11	5	5	23.8	20.8
Severe impairment	57 (55)	20	15	22	38.6	75.9
1st follow-up to 2nd follow-up	104 (100)	51 (49)	24 (23)	29 (28)		
No impairment	38 (37)	34	3	1	89.5	66.7
Moderate impairment	23 (22)	10	11	2	47.8	45.8
Severe impairment	43 (41)	7	10	26	60.5	89.7

^a CPT performance grouping: No impairment (NI): adjusted z score of $d' \ge -1$; Moderate impairment (MI): $-2.5 \le$ adjusted z score of d' < -1; Severe impairment (SI): adjusted z score of d' < -2.5.

Table 4 Variables associated with the group assignment to severe impairment for the undegraded CPT performance at various time frames from multiple logistic regression analysis a (n = 104)

Time frame for the outcome	Predictors	β (SE)	p	Odds ratio (95% CI)
Baseline to 1st follow-up				
$(r^2 = 0.46)$	Previous CPT d' group	0.81 (0.38)	0.03	2.24 (1.54-3.29)
	Concurrent CPT $\ln \beta$	0.75 (0.23)	0.001	2.13 (1.68–2.66)
1st follow-up to 2nd follow-up				
$(r^2 = 0.64)$	Previous CPT d' group	2.20 (0.63)	< 0.001	9.01 (4.81–16.95)
	Concurrent CPT $\ln \beta$	0.75 (0.25)	0.003	2.71 (2.23–2.72)
Baseline to 2nd follow-up				
$(r^2 = 0.49)$	Previous CPT d' group	0.99 (0.43)	0.02	2.68 (1.75-4.14)
	Concurrent CPT $\ln \beta$	1.07 (0.26)	< 0.001	2.91 (2.25–3.78)

^a The follow-up CPT group assignments were regressed on previous CPT group assignments with covariates including duration of illness, age at onset, interval between follow-ups, antipsychotic dosage, concurrent CPT $\ln \beta$, PANSS negative, disorganization, positive, depressive, and excitement factor scores. Only covariates with significant regression coefficients (p < 0.05) were shown.

subjects with severe impairments in attention remained so through out the period. The correlations in the CPT performances across assessment time points were "good" but not "excellent" in terms of test-retest reliability, indicating that CPT performances changed considerably during a long follow-up period. The tendency to improve in the whole sample was supported by the data-driven growth mixture modeling analysis, in which three major subgroups could be demarcated by their trajectories and initial levels of impairments, as well as the subsequent subgrouping analyses. The patterns of shifting between subgroups might partly arise from the fact that the NI and SI subgroups changed subgroup status less than the middle group simply because that they each had only one direction in which to move, while patients in the MI subgroup could move either up or down in the distribution. However, the unanimous direction toward improvements made such an explanation less likely to account fully for the findings.

At first glance, the results did not seem to support CPT deficits as stable vulnerability indicators. However, it was notable that the proportion of patients with low perfor-

mance over the follow-up period was 27.1% in the growth mixture modeling, which was comparable to the figure of 24.1% for those with severe impairments at baseline that remained severely impaired throughout the follow-up period. Moreover, the most important determinants of later poor CPT performances were found to be previous severe impairments, implying that CPT deficits with sufficient severity were not substantially influenced by extraneous factors and sufficed to predict their own further development even after long follow-up intervals. In this regard, severe CPT deficits probably reflected the core neurobiological vulnerability traits.

Some precaution must be pointed out, however. Deficits across multiple neurocognitive domains have been considered characteristic of schizophrenia (Flashman and Green, 2004). Their associations with obstetric complications (Cannon et al., 2002), structural brain changes and subtle neurological signs (Arango et al., 2000; Leask et al., 2002; Obiols et al., 1999) suggested probable neurodevelopmental origins (Cannon et al., 2003; Lawrie et al., 2001; Lewis and Levitt, 2002). In this regard, the observed persis-

tent deficits on CPT might be one part of the static encephalopathy associated with early developmental insults (Hill et al., 2004). In addition, later disease- and treatment-related neurocognitive side effects might have also contributed to the sustained attention deficits. Indeed, when attention deficits alone were employed to predict the occurrence of schizophrenic spectrum disorders, unacceptable high false positive rate ensued (Keri and Janka, 2004), underscoring the part played by non-genetic factors. However, studies among nonpsychotic relatives of schizophrenia patients (Chen et al., 2004; Ross et al., 2000) found that the neurocognitive deficits per se were familial in these families. Our recent finding that siblings from multiplex families exhibited worse CPT performance than did their counterparts from simplex families (Tsuang et al., in press) provided further support for the proposition that impairment in sustained attention is a genetic vulnerability marker for schizophrenia. Further strategy to disentangle genetic from non-genetic factors on CPT deficits thus should focus on the dosage effect in the transmission pattern within families of those with persistent severe impairments, in contrast with that within less impaired patients' families. In addition, other signs of early brain injury and neurological dysfunction should be carefully sought and controlled for (Lawrie et al., 2001; Leask et al., 2002; Lewis and Levitt, 2002).

The substantial proportion of subjects with CPT performance within population norm was another distinct subgroup and could be considered as lying at the 'normal' end of the liability continuum. Neuropsychologically normal schizophrenia subjects have been a non-rarity and they demonstrated a clinical profile and disease course different from those with abnormal performances (Kremen et al., 2000; Palmer et al., 1997). Although there might still be some decline from premorbid level of performance (Kremen et al., 2000), these patients were nonetheless considered relatively, though not absolutely, devoid of the disease/genetic vulnerability-related decrements. In this regard, they could serve as internal contrast group within schizophrenia proper to disentangle the effects of clinical state-related and genetic-related factors, hence a powerful tool to explore the issue of genetic heterogeneity within schizophrenia.

In contrast to previous two subgroups, those with CPT performance at intermediate level showed marked instability during the long-term course, suggesting that their performances were mainly state-dependent. In this regard, the prospect that some major gene effects be found from this group is expected to be meager. In addition, since many recent pharmacological interventional studies on sustained attention deficits were carried out in patients with modest degree of impairments with short-term follow-up (Cornblatt et al., 1997; Epstein et al., 1996; Hong et al., 2002), the intrinsic instability of this group's short-term performances rendered the results as preliminary at best.

In logistic regression analyses, only two factors consistently emerged as significant predictors for CPT performance group assignment, i.e., previous CPT grouping status and concurrent response criterion $\ln \beta$, whereas variables of disease course and clinical symptoms did not. The results seemed to be inconsistent with previous findings that certain clinical symptom dimensions were associated with the CPT deficits, especially negative and thought/disorganization symptoms (Guillem et al., 2001; Liu et al., 1997; Mass et al., 2000; Nuechterlein et al., 1986; Stratta et al., 2000; Strauss, 1993). Several reasons might have accounted for this discrepancy. First, meta-analysis on the associations between CPT and psychopathology revealed low to moderate associations (weighted r = -0.31), with negative symptoms accounting for less than 10% of the variances in CPT d' (Nieuwenstein et al., 2001). Second, previous studies either were limited to cross-sectional analyses (Mass et al., 2000; Stratta et al., 2000) or examined only cross-temporal correlations between the two in a short period of time (Nuechterlein and Dawson, 1984). Although both sustained attention deficits and negative symptoms were enduring features of schizophrenia, it is still possible that they were determined by distinct pathophysiological mechanisms rather than a common underlying neurocognitive deficit (Tamminga et al., 1998). In this regard, the reported associations between negative symptoms and sustained attention deficits might not be an invariable phenomenon for all schizophrenia patients.

It was also noteworthy in current study that CPT performances were significantly influenced by the response strategy. Our past studies revealed that schizophrenia patients seemed to take a more lenient strategy when undertook a more difficult version of the CPT. In contrast, when an easier attentional task was administered, patients were less likely to disengage their attention and little effects on performances would be expected (Liu et al., 2000; Liu et al., 2002; Liu et al., 1997). In addition, $\ln \beta$ was found to be independent from the sensitivity index d' in community subjects, but the two indices exhibited a modest correlation in schizophrenia patients (Chen et al., 1998a). Thus, schizophrenia patients' deficits in CPT performances should be carefully examined in terms of the task-taking strategy, especially when more difficult versions with perceptual loads were used.

Limitations of this study should be kept in mind while interpreting the results. First, the lack of a longitudinal normal control group in this study made it difficult to determine whether the subgroup shifting could be found in normal subjects too as a result of natural tendency for fluctuations in CPT performances. However, there has been evidence that performances in multiple neurocognitive domains, in which attention was included, remained stable across assessments in normal subjects, hence not likely to account for the changes observed in schizophrenia patients (Heaton et al., 2001). Second, subjects included for analysis 1 and analysis 2 were different subpopulations of the total sample and the results might have been systemically biased. Nevertheless, comparisons in the demographic

and clinical characteristics did not reveal significant differences between those included and not included for both analyses 1 and 2. Compatible results revealed by both analysis 1 and analysis 2 further indicated that the selection of subjects for analysis 2 might not have severely distorted the results. On the other hand, those without baseline CPT data were mainly from those unable to complete even the undegraded version, hence more likely to be severely impaired. The proportion of the SI subgroup thus might have been under-estimated. Hence, the associations between CPT deficits and negative symptoms might have been underestimated in this study.

In summary, this study demonstrates that there were distinct subgroups in schizophrenia patients that exhibited different trajectories in the long term course of sustained attention deficits, which might have important theoretical and practical repercussions for further research into the heterogeneity of schizophrenia. Despite the fluctuations, there is indeed a subgroup of severely impaired patients whose sustained attention deficits might represent a stable vulnerability indicator for schizophrenia.

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