Original Article

New information on age composition and length-weight relationship of bluefin tuna, *Thunnus thynnus*, in the southwestern North Pacific

Chien-Chung HSU, 1 Hsi-Chiang LIU, 1,* Chi-Lun WU, 2 Shih-Tsung HUANG AND Hsueh-Keng LIAO 2

¹Institute of Oceanography, National Taiwan University, PO Box 23-13, Taipei, Taiwan 106 and ²Taiwan Fisheries Research Institute, Keelung, Taiwan 202

SUMMARY: New information on catch-at-age composition and length–weight relationships is presented for Pacific northern bluefin tuna caught by Taiwanese small-scale longliners in the southwestern North Pacific. The fork length–eviscerated weight relationship of Pacific northern bluefin tuna, *Thunnus thynnus*, caught off southwestern North Pacific (off the Taiwan coast) was determined and compared with relationships previously reported by various studies for different waters. The best representative of this relationship is $W=0.000023058L^{2.9342}$, where W is the eviscerated weight (kg) and L the fork length (cm). The eviscerated weights estimated from this relationship were multiplied by a scaling factor of 1.112 to obtain the estimated round weight. The visual comparison of weightat-length showed that there were almost no differences between the relationships previously reported for the northwestern Pacific, the Sea of Japan and the present study. This indicated that Pacific northern bluefin tuna from those three waters could be of the same stock. Further assessment studies should combine information from those waters.

KEY WORDS: length–weight relationship, likelihood ratio test, northern bluefin tuna, *Thunnus thynnus*.

INTRODUCTION

Northern bluefin tuna, *Thunnus thynnus*, are found in the northern hemisphere, but very few are found in the southern hemisphere.¹ In worldwide terms, three bluefin tuna stocks are assumed for assessment and management purposes, those are the western Atlantic stock, the eastern Atlantic and Mediterranean stock and the Pacific stock.

Since the 1970s, the International Commission for the Conservation of Atlantic Tunas (ICCAT) have conducted many studies on the two Atlantic stocks.² Those studies have been published in the ICCAT Collective Volume of Scientific Papers. However, few studies have been done on the Pacific stock.³⁻⁷ It is noted that the Atlantic stocks have been assessed as over-exploited and have been managed by size limit, time and area closures and catch quota renewal year by year.² The stock status of the Pacific bluefin tuna is still unknown, due to poor knowledge of the biological parameters and statistics used for evaluating the stock.

Analytical assessment studies generally require the determination of the number of fish caught at suitable length and the length-weight conversion in the commercial fisheries.^{8,9} In many cases, this is achieved by the adjustment of length frequencies obtained by port sampling to the total reported catch in weight through the application of length-weight relationship. A formulation of the length-weight relationship is therefore essential not only in those assessment techniques, but in catch compilation from number into weight. Although many fork length-weight (eviscerated and round) relationships of eastern and western North Pacific bluefin tuna have been formulated for internal use by Japanese and American scientists, no length-weight relationship was provided for bluefin tuna from southwestern North Pacific, especially off the Taiwan coast.¹⁰ This is because weighing very large and valuable bluefin tuna on board is difficult, and most fish landed are eviscerated (gilled and gutted with intact head), which results in the difficulty of accurately measuring round weight. However, a fork length-weight relationship (LW) should give valu-

^{*}Corresponding author: Tel: 886-2-2362-2987. Fax: 886-2-2366-1198. Email: hsucc@ccms.ntu.edu.tw

Received 3 August 1999. Accepted 2 February 2000.

able insights both into the most appropriate catch conversion between number and weight and management units.

Therefore, the objectives of this study are to analyze the age-class variation derived from size composition, and to formulate the length–weight relationship in order to contribute new useful information for the stock assessment of Pacific northern bluefin tuna stock caught off the Taiwan coast.

MATERIALS AND METHODS

Data collection

Two kinds of data, length only and length and weight combined, were collected for Pacific northern bluefin tuna from the southwestern North Pacific (off the Taiwan coast) (Fig. 1) during the fishing seasons from 1993 to 1998. All length measurements were made as fork length in the present study.

Fork lengths (FL) were measured for almost all the catches from 1993 to 1998. However, fork lengths with eviscerated weight of bluefin tuna were collected only during 1997–1998 from Taiwanese longline fleets operating in the waters of the southwestern North Pacific (Fig. 1). In addition to these samples, a significant amount of small bluefin tuna (FL <170 cm) was measured and weighed at various times.

In Taiwan, the small-scale longline fleets (<100 gross register tons, GRT) operate offshore in the southwestern North Pacific near the Taiwan coast throughout the year, and they target various tunas, mainly yellowfin tuna



Fig. 1 Location of the fishing area for Taiwanese small-scale longline fleets targeting Pacific northern bluefin tuna. The two major landing ports, Tungkang and Suao, for unloading Pacific northern bluefin tuna are also shown.

(*Thunnus albacares*), bigeye tuna (*T. obesus*) and albacore (*T. alalunga*). A trip is about 7 days on average. Nonetheless, some of the fleets operate to target bluefin tuna during April and June each year. Other fisheries have never targeted bluefin tuna until recently. Tungkang, located at southwestern Taiwan, is the largest landing port and Suao at northeastern Taiwan is the next biggest (Fig. 1). The fishes caught were chilled in ice, and most bluefin tuna were gilled and gutted on board to maintain the meat quality. All the processed bluefin tuna were then exported directly to Japan by air immediately they were unloaded at those two ports. For the present study, some of the longline fleets were asked for weight measurements as they unloaded.

The fish were then measured at landing ports (Tungkang and Suao in Fig. 1) as they were unloaded. The fork length (from the tip of snout to tail fork) was measured by steel calipers to the nearest cm, and the weight was determined to the nearest kg by scales (a Chinese steelyard, Nankuang Measurement Instruments, Inc., Taipei, Taiwan). Only fork length was measured for over 90% of the gilled and gutted samples. The sexes were not distinguished because the fish had been eviscerated on board.

Size/age composition

Because fork lengths of almost all bluefin tuna caught were measured at landing ports, the length frequencies may be sufficient to represent the size distributions of the studied population. If the normality of actual size distribution was tested for having a similar dispersion from 1993 to 1998 before any statistical comparison was made. The normal distribution of annual actual size distribution was examined by g^1 and g^2 statistics for its skewness and kurtosis, respectively.¹¹ Then, to investigate the changes of major age-class in the annual catch, the difference of age compositions of the annual catch was examined by separating the catch-at-length based on the growth equation: $FL = 320.5(1-e^{-0.1305t-0.0728})$ into catch-at-age.¹²

Statistical analyses

Examining scatter plots of the LW relationship, a power function, $W = \alpha L^{\beta}$, where W is the weight (kg) and L is the fork length (cm), seems appropriate for deriving the LW relationship. Thus, least squares linear regressions of eviscerated weight on fork length were derived for samples after natural logarithmic transformation of the two variables.

To test the equality that samples are collected from the same population, the LW relationships were first derived by monthly samples. The likelihood ratio test was carried out to test the null hypothesis that the estimated parameters are constant across the monthly sample data sets with different variances.^{12,13} The test was made on combining the slope and the intercept together with a likelihood ratio test.¹³

Test of equality of parameters between data sets based on the theory of likelihood ratio tests have been introduced by several scientists (*e.g.* Kimura (1980)¹⁴; Cerrato (1990)¹⁵ and Quinn and Deriso (1999).¹³ The likelihood ratio tests can be used to compare full models with a reduced model for two or more data sets. By assuming a normal distribution with additive errors for five data sets, the maximum likelihood estimation (MLE_i) of parameters $\hat{\Theta}_i$, and standard deviation $\hat{\sigma}_i$, for data set *i*, Y_i with sample sizes, n_i results in:

$$\max \ln L_i(\hat{\Theta}_i, \hat{\sigma}_i | \{Y_i\}) = -\frac{n_i}{2} [\ln(2\pi\hat{\sigma}_i^2) + 1], \quad (1)$$

and

$$\hat{\sigma}_{i}^{2} = \sum_{j=1}^{n_{i}} \left(Y_{ij} - \hat{Y}_{i} \right)^{2} / n_{i}.$$
(2)

The joint maximum log likelihood $\ln L_F$ for the full model is obtained by:

$$\ln L_F = \sum_{i=1}^R \max \ln L_i.$$
(3)

Meanwhile, the maximum log likelihood for reduced model, $\ln L_R$, is then from (3) with n_i and $\hat{\sigma}_i^2$ replaced by n and $\hat{\sigma}^2$. Therefore, the likelihood ratio test statistics is:

$$\chi^2 = -2\ln\left(\frac{L_R}{L_F}\right) \tag{4}$$

The asymptotic distribution is a chi-square distribution with degrees of freedom equal to the distribution in the degrees of freedom between the full model and reduced model.

The hypothesis test is $H_0: \Theta_i = \Theta_j$ for all pairs (i,j) vs $H_a: \Theta_i \neq \Theta_j$ for at least one pair (i,j). There are f = Rp - p = (R-1)p degrees of freedom, where R is the number of data sets and p is the number of parameters. A significant ratio indicates a reduced model is not statistically similar to the full model.

If the equality of those LW relationships obtained from different temporal scales was satisfied, all the samples were combined to derive the LW relationship for the whole samples. The resultant LW relationship was compared graphically with other studies.¹⁰ However, statistical analyses were not applied to the comparison because only mean values were available for all of the other studies.

RESULTS

The annual length compositions (Fig. 2) of Pacific northern bluefin tuna taken by Taiwanese longline fleets were examined for normality. The results of the normal-

ity examination for the 1993–1998 annual actual size distributions show either skewness or kurtosis to depart from a normal distribution: skewness to the right and leptokurtosis in 1993, 1997 and 1998, and platy-kurtosis in 1994, 1995 and 1996 catch-at-size distributions. Hence, the normality test was not met so that the test of dispersion was not necessary for size composition examination, because there were different age-classes caught from year to year.

Furthermore, the catch-at-size (Fig. 3), which was compiled from annual actual size composition, was separated out using the Yukinawa and Yabuta (1967) growth equation.¹² As previously reported, the reproduction period for the study species is during June.¹⁶ Then, the age was taking account from July to June of the next year instead of from January to December as is usual. Therefore, the catch-at-age was obtained as shown in Fig. 3, revealing that the major age-classes of Pacific northern bluefin tuna caught by Taiwanese small-scale longline fleets varied from year to year. The major age-classes are 8-year-old in 1993, 6-year-old in 1994, 6- and 7-year-old in 1997 and 7-year-old in 1998.

Five sets of LW measurements were made on 1774 Pacific northern bluefin tuna collected off the eastern Taiwan coast from May and June, 1997 and 1998, and one set on 318 samples collected irregularly at various times. The latter samples have most fork lengths smaller than 170 cm, but the former samples ranged between 170 and 290 cm. Due to the different sampling time period, five data sets sampled were separately fitted by the power function, $W = \alpha F L^{\beta}$. The likelihood ratio test was used to examine the equality of parameters of the LW relationships derived from those different data sets.

All of the derived LW relationships are highly significant (P < 0.001) among five regression lines (Table 1). The residual analysis (Fig. 4) illustrates those current fits are appropriate, assuming that the normal distribution with an additive error for the LW relationship was met. Furthermore, to take account of two parameters together for the LW relationship model, the likelihood ratio test was pursued as Eqns 1–4 for comparing five LW relationships. The result of likelihood ratio test (Table 2) indicates no significant difference (P < 0.001) among the five samples collected. It may indicate that those five LW relationships are identical in their tendency and those five data sets of the samples could be combined for the derivation of a LW relationship for Pacific northern bluefin tuna.

The relationship between fork length (*L*) and eviscerated weight (*W*) could be derived from 1774 combined individuals (Fig. 5). The relationship yielded an R^2 of 0.987 (r=0.993, P<0.001):

$W = 0.000023058L^{2.9342}$.

Comparison of the LW relationship derived here with those from various studies (Table 3) suggested very little



Fig. 2 Fork length distribution of Pacific northern bluefin tuna caught by Taiwan small-scale longline fleets operating in the southwestern North Pacific (off the Taiwan coast) from 1993 to 1998, where N is the individuals measured, x is the mean and SD is the standard deviation, g_1 and g_2 statistics are used to examine the skewness and kurtosis of the normal distribution, respectively.

Table 1 Regression parameters and statistics for the fork length (*FL*, cm)–eviscerated weight (*W*, kg) relationships, $W = \alpha FL^{\beta}$ of Pacific northern bluefin tuna collected from the different seasons of May and June, 1997 and 1998, respectively, and the small-sized specimen from various times before 1993

| Fishing seasons | Ln – Ln regression | | | | | | | | |
|-----------------|--------------------|---------------|----------------|-----|--------|-------------------------|----------------|-----------|-------|
| | Σ_x^2 | Σ_{xy} | ${\Sigma_y}^2$ | n | β | $\alpha \times 10^{-5}$ | Residual SS | F | r |
| May 1997 | 1.9267 | 5.2794 | 16.4306 | 313 | 2.7401 | 6.6400 | 1.9644 | 1145.13** | 0.938 |
| June 1997 | 2.6666 | 8.0564 | 26.0430 | 372 | 3.0212 | 1.3960 | 1.7028 | 2644.43** | 0.967 |
| May 1998 | 1.4871 | 4.3009 | 14.1481 | 431 | 2.8922 | 2.9899 | 1.7093 | 1560.94** | 0.938 |
| June 1998 | 1.5954 | 4.6505 | 15.3099 | 340 | 2.9149 | 2.5174 | 1.7540 | 1306.14** | 0.941 |
| Small-sized | 8.4819 | 24.3450 | 72.3274 | 318 | 2.8702 | 3.1206 | 2.4511 | 4504.22** | 0.983 |

r is the correlation coefficient and ** is significant at P < 0.0001.



Fig. 3 The age composition of the catches of Pacific northern bluefin tuna by Taiwanese small-scale longline boats in the southwestern North Pacific waters from 1993 to 1998, where the catch-at-age was sliced from catch-at-length raised to catch with the growth equation of Yukinawa and Yabuta (1967).¹²

or no difference from that reported by Ishizuka (1989)³ and Anon. (1995),¹⁰ but different to Anon. (1990)¹⁷ (Fig. 6). Ishizuka (1989)³ and Anon. (1995)¹⁰ also used eviscerated weights to estimate the LW relationship for Pacific northern bluefin tuna from the northwestern

Pacific, and Anon (1995)¹⁰ used round weights of northern bluefin tuna from the Atlantic. The LW relationship collected off the Sea of Japan,¹⁰ northwestern Pacific waters³ and off the Taiwan coast (the present study) shows that there is no significant difference in visual



Fig. 4 The residual plot of fitting length–weight relationships derived from five data sets for assuming a normal additive error, in which (a) is from the data sample in May 1997, (b) in June 1997, (c) in May 1998, (d) in June 1998 and (e) the small-sized fish. Results show that residual distributions seem coincident to the error assumption.



Fig. 5 The fork length (L) – eviscerated weight (W) relationship of Pacific northern bluefin tuna collected from southwestern North Pacific (off Taiwan Coast).



Fig. 6 Comparisons of the length–weight relationships derived from the present study with some previously reported curves, where PW is the eviscerated weight (kg), W is the round weight (kg) and L is the fork length (cm). The curves correspond to equations from the top down, and curves numbered 1 and 2, 3 and 5 are too similar to be distinguishable.

Table 2The likelihood ratio tests for the parameter equality of the length–weight relationships of Pacific northern bluefin tunaderived from five different data sets collected from the southwestern North Pacific

| Statistics | Specimen | | | | | |
|-----------------------------------------------------------------------------------------------------------------------------|----------|-------------|-----------|----------|-----------|-----------------------------------------------------|
| | May 1997 | Small-sized | June 1997 | May 1998 | June 1998 | |
| $ \frac{\max \ln L_i \text{ (Full)}}{\max \ln L \text{ (Reduced)}} $ $ \chi^2_{\text{critical}}, d.f. = 8, P = 0.01 $ $ P $ | 349.4859 | 322.3914 | 474.0657 | 580.1572 | 412.9584 | 2139.0586 2113.9798 0.0236 20.09 0.0001 |

The result shows no significant difference for those length-weight relationships derived from five different data sets.

| Length–weight relationship | Length ranges (cm) | Stock | Authors | |
|-------------------------------------------|--------------------|-------------------|------------------------|--|
| $W = 1.961 \times 10^{-5} L^{3.0092}$ | Not available | Mediterranean Sea | Anon., 1990 | |
| $W = 1.9 \times 10^{-5} L^{3.0}$ | <100 | Eastern Atlantic | Cited from Anon., 1995 | |
| $W = 4.044 \times 10^{-5} L^{2.837}$ | >100 | Western Atlantic | Cited from Anon., 1995 | |
| $PW = 4.073 \times 10^{-5} L^{2.8344}$ | >99 | Western Pacific | Ishizuka, 1989 | |
| $PW = 1.982 \times 10^{-5} L^{2.9677}$ | 60–264 | Japan waters | Anon., 1995 | |
| $PW = 2.3058 \times 10^{-5} \ L^{2.9342}$ | 50–290 | Taiwan waters | Present study | |
| | | | | |

 Table 3
 Some representatives of the length–weight relationship of northern bluefin tuna

(L), the length used is fork length; PW, indicates the eviscerated weight that was measured after gilled and gutted processing; and W, the round weight.

comparison, and the similar LW relationship may suggest that these data sets were sampled from the same stock. However, the result obtained from the current study provides samples from a wider range of fork length to derive the LW relationship, and provides a LW relationship of Pacific northern bluefin tuna from the southwestern North Pacific region that is different from the traditional areas of the northwestern North Pacific.

DISCUSSION

The current study contributes two important features for further assessment on Pacific northern bluefin tuna.

First, the variation of age-classes derived from the annual catch of Taiwanese small-scale longliners may reveal that the presence of the species around Taiwan waters is on route for spawning migration. The exact migratory route for this species is not very clear, but the two assumed spawning grounds are in the Sea of Japan and around the waters off Yaeyama Islands. The latter spawning aggregation was fished by tuna longliners.^{16,18} Also the progressive larvae surveys show that sizes of larvae of Pacific northern bluefin tuna are smaller caught from the southwestern than the northwestern North Pacific (TS Chiu, Department of Zoology, National Taiwan University, pers. comm. 1996). This survey may indicate that the species spawn somewhere in the waters off eastern and/or northeastern Taiwan and the larvae gradually drift northward to the waters around Japan by the kuroshio current. Therefore the age composition obtained in the current study may partially support this prediction, because almost all the Taiwan catches are mature adults fished from the southwestern North Pacific. To provide evidence for this prediction, a spatial and temporal study on ovary development is currently being conducted by our staff.

Second, the current study provides a new LW relationship derived from a wider fork length range (FL = 50-290 cm; Table 3) than previously reported. The equation could be used in catch conversion from number into

weight and in the assessment analyses, if the stock in southwestern North Pacific caught by Taiwanese smallscale longliners is from similar stock to those from the waters around Japan caught by Japanese fisheries. Reviewing the size/age composition, very large fish were rare in the catches of northwestern North Pacific,⁷ but this was also found in the present study. The adult fish distribute in the southwestern North Pacific and Nansei Shoto area.¹⁶ Traditionally, two stocks were assumed for informally managing the study species, but a preliminary study using micro-satellite loci showed no difference among the samples from the North Pacific.¹⁹ If the Pacific northern bluefin tuna belong to one stock, the current LW relationship result will be useful in future assessment work. To consider all Pacific northern bluefin tuna as a unit stock from the western North Pacific is essential as large- and small-sized fish can be pooled together in an analysis. Accordingly, it seems insufficient that only the LW relationship comparison be used to indirectly depict the possibility of classifying stock. Also from a biological point of view, it would be possible to classify a stock as similar to the northwestern Pacific because of the actual small differences in weight at length.^{3,10} Based on management purposes, two stocks were primarily assumed; however, we suggest the biological viewpoint needs to be created in future to identify the stock structure.

Although there were differences in time for sampling data sets used to derive the LW relationship, the result obtained by likelihood ratio test shows no statistical significance (P < 0.0001) for the model used. Thus, it is reasonable to combine five data sets to estimate the LW relationship. However, robust regression procedures, such as the least median of squares were used to diagnose outliers within the data sets.^{20,21} It is difficult to identify outliers from a residual plot, as shown in Fig. 4, because the outliers always pull the regression line towards themselves.²² Thus the outlier detection was made for some regression analyses.^{23,24} However, we have not proceeded with this procedure because the likelihood ratio test shows no difference among data sets (Table 2) and the outlier detection may not always make the LW relationship fitting different.²⁴

Accurately weighing very large northern bluefin tuna caught by longline fleets is difficult; however, counting individuals caught and measuring fork lengths are possible. Therefore, the round weight-at-length can be estimated from this LW relationship multiplying a scaling factor of 1.112.10 Moreover, the widest range of fork lengths was used to derive the LW relationship in the present study (cf. Table 3). As usual, the annual catch of Taiwan longline fleets was based on landings reported by trade agencies and logbooks recorded in numbers and submitted by captains.⁸ There may be under-reporting in the trade reports, and inaccuracy can exist with on-board weighing, especially for this very large and valuable northern bluefin tuna. Hence, the LW relationship derived here is thought to be a useful tool for estimating the numbers of fish caught at length or age, and for compiling the catch from individuals to total biomass in the Taiwan commercial longline fisheries.

Stock assessments for Pacific northern bluefin tuna were not available. Although Nishida⁷ and Watters²⁵ have derived standardized catch rates of Pacific northern bluefin tuna, actual implementations were not begun due to lack of catch data integration, biological characteristic parameters estimation and fundamental fisheries knowledge. In Taiwan, the Pacific northern bluefin tuna fishery began in the early 1950s using small-scale longline fleets (<100 GRT), but became seasonally targeted from 1990 during April-June each year. The annual catch of the species was according to the migratory age-classes off the Taiwan coast. To date, Pacific northern bluefin tuna is becoming one of the most economically important species for Taiwan smallscale longline fleets. Therefore, fishery-dependent data collection and biological studies are being pursued. Moreover, further virtual population analysis assessment is in progress for Pacific northern bluefin tuna using the LW relationship derived here and size compositions collected in the present study.

ACKNOWLEDGMENTS

This study was fully supported by the Council of Agriculture, Taiwan in contract no. 88-AST-1.4-FID-06(06). We also thank the critical comments of two anonymous referees for their valuable advice.

REFERENCES

- Smith PJ, Conroy AM, Taylor P. Biochemical-genetic identification of northern bluefin tuna *Thunnus thynnus* in the New Zealand fishery. NZ. J. Mar. Freshwater Res. 1994; 28: 113–118.
- Anon. Detailed report bluefin tuna. International Commission for the Conservation of Atlantic Tunas. In: Collective Volume of Scientific Papers Madrid, Spain. 1999.

- Ishizuka Y. Estimates of Catch and Age Compositions for Northern Bluefin Tuna (*Thunnus Thynnus*) Caught by Japanese Fisheries in the Pacific Ocean During 1966 to 1986. National Research Institute of Far Seas Fisheries, Shimizu, Shizuoka, Japan. 1989.
- Bayliff WH. Status of northern bluefin tuna in the Pacific Ocean. Inter-American Tropical Tuna Commission, Special Report 1991; 7: 29–88.
- Bayliff WH. A review of the biology and fisheries for northern bluefin tuna, *Thunnus thynnus*, in the Pacific Ocean. FAO Fisheries Report 1994; 336: 244–295.
- Bayliff WH, Ishizuka Y, Deriso RB. Growth, movement, and attrition of northern bluefin tuna, *Thunnus thynnus*, in the Pacific Ocean, as determined by tagging. *Inter-Am. Trop. Tuna Comm. Bull.* 1991; 20: 1–94.
- Nishida T. Preliminary attempt to standardize catch rate of Pacific northern bluefin tuna (*Thunnus thynnus*) based on the Japanese longline fisheries data. Paper prepared for the Pacific Northern Bluefin Tuna Workshop, held in the Inter-American Tropical Tuna Commission, La Jolla, CA. 1997; May 23–24.
- Hsu CC, Lin MC. The recent catch estimating procedures of Taiwanese longline fisheries. International Commission for the Conservation of Atlantic Tunas. *Collective Volume of Scientific Papers* 1996; 43: 175–177.
- 9. Hsu CC. The length-weight relationship of albacore, *Thunnus alalunga*, from the Indian Ocean. *Fish. Res.* 1999; **41**: 87–92.
- Anon. Review of length–weight and age-length relationship on Pacific northern bluefin tuna. In: *Report on Pacific Northern Bluefin Tuna Survey Around Japan*. National Research Institute of Far Seas Fisheries, Shimizu, Shizuoka, Japan. 1995; 371–386 (in Japanese).
- Yukinawa M, Yabuta Y. Age and growth of bluefin tuna, *Thunnus thynnus* (Linnaeus), in the north Pacific Ocean. *Nankai Regional Fish. Res. Lab. Rep.* 1967; 25: 1–18.
- Sokal RR, Rohlf FJ. Biometry, 3rd edn. W. H. Freeman and Company, New York. 1995.
- Quinn TJ, Deriso RB. Quantitative Fish Dynamics. Oxford University Press, New York. 1999.
- Kimura DK. Likelihood models for the von Bertalanffy growth curve. Fish. Bull. 1980; 77: 765–776.
- Cerrato RM. Interpretable statistical tests for growth comparisons using parameters in the von Bertalanffy equation. *Can. J. Fish. Aquat. Sci.* 1990; 47: 1416–1426.
- Tanaka S. Spawning season of Pacific bluefin tuna estimated from histological examination of the ovary. Manuscript submitted to NRIFSF/IATTC Joint Workshop on Pacific Northern Bluefin Tuna. December 8–9, 1999. Shimizu, Japan. 1999.
- Anon. Report of GFCM/ICCAT Expert Consultation on Evaluation Stocks of Large Pelagic Fishes in the Mediterranean Area. International Commission for the Conservation of Atlantic Tunas. Collective Volume of Scientific Papers 1990; 33: 44.
- Shimizu H, Yano K. The fishing of bluefin tuna in spawning aggregations around Yaeyama Islands. Paper submitted to the NRIFSF/IATTC Joint Workshop on Pacific Northern Bluefin Tuna, December 8–9, 1999, Shimizu, Japan. 1999.
- Tagaki M, Okamura T, Chow S, Taniguchi N. PCR primers for microsatellite loci in tuna species of the genus *Thunnus* and its application for population genetic study. *Fisheries Sci.* 1999; 65: 571–576.
- Rousseeuw PJ. Least median of squares regression. J. Am. Stat. Assoc. 1984; 79: 871–880.
- Rousseeuw PJ, Leory AM. Robust Regression and Outlier Detection. Wiley, New York. 1987.

- 22. Sen AK, Srivastava MS. Regression Analysis: Theory, Methods and Applications. Springer, New York. 1990.
- Chen Y, Jackson DA, Paloheimo JE. Robust regression approach to analyzing fisheries data. Can. J. Fish. Aquat. Sci. 1994; 51: 1420–1429.
- 24. Hsu CC. The length-weight relationship of albacore, *Thunnus alalunga*, from the Indian Ocean. *Fisheries Res.* 1999; **41**: 87–92.
- 25. Watters G. CPUEs of Pacific bluefin tuna in a core fishing area extending from the coast of Japan to 150°W: using regression trees to hypothesize about CPUEs times and areas in which there was no fishing. Paper submitted to the NRIFSF/IATTC Joint Workshop on Pacific Northern Bluefin Tuna, December 8–9, 1999, Shimizu, Japan. 1999.